

Lec 5a Some historical review.

So far we have seen some brief historical review, and a couple of important theorems:

1) Subadditivity, Fekete's lemma

2) Existence of time constant

3) Simple comparison inequalities for the time constant.

4) Jorge will show us the subadditive ergodic theorem.

Now I want to return to the history, starting with

(1961) Eden. He wants to study Morphogenesis

(morpho-"form", genesis-"creation")

You have an individual cell. It has some

"information" inside it, say its DNA code

(by then genes (1944) and the existence of DNA

(Watson, Crick, Wilkins and Franklin) (1950s).

X-ray crystallography

So this cell divides and it is only influenced by its random environment of nutrients and other cells. How does it achieve its final shape? How the heck does a single brain cell divide and form a brain? Amazing, right?

Eden wanted to study this question, asked by Alan Turing (1952) ("information theory") using a Toy model.

Let us read the intro starting in paragraph 2.

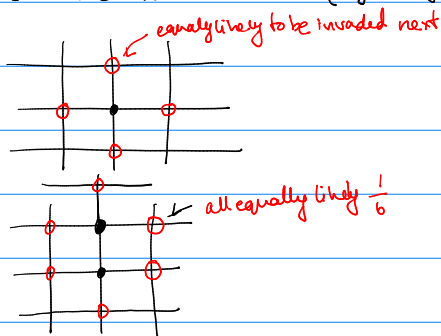
- Even twins are not exactly identical, so must be influenced by environment.

(Fingerprints of twins, armadillo scale counts, skin color patterns of cows)

More influence of environment: growth of chicken cells into a kidney tubule or a feather.

Turing's model was 1D (Does someone want to read it)

Eden's model was in 2D. (Pg 9 of his paper)



Biological counterpart: 2D growth of cells, where growth only happens at the periphery.

- *Ulva lactuca*, common sea lettuce.
- *Pennisetum setaceum*.

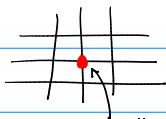
On page 236 (12) there are some pictures he generated. Looks pretty cool right? He claims that it's a ball.

It's worth reading his paper in some detail.

Richardson (1973)

He was inspired by Williams-Bjerknes process. On

the square lattice.



"The hazard

function" for ^{the time to} cell division is constant.

When cell division occurs, it picks a neighbor at random. If this neighbor is white, it will turn red.

$$\text{Hazard function: } h(t) = \frac{f(t)}{\int_t^{\infty} f(y) dy} = \frac{\overset{\text{random death times in } [t, t+\Delta]}{P(T=t)}}{P(T>t)}$$

The only distribution with constant hazard for is

$h(y)$ = "conditional density of dying now given that you have survived up to now."

$$f(y) = \lim_{\Delta \rightarrow 0} \frac{F(y+\Delta) - F(y)}{\Delta}$$

$$h(y) = \lim_{\Delta \rightarrow 0} \frac{F(y+\Delta) - F(y)}{\Delta P(T>y)} = \lim_{\Delta \rightarrow 0} \frac{P(y < T \leq y+\Delta)}{\Delta P(T>y)}$$

$$= \lim_{\Delta \rightarrow 0} \frac{P(y < T \leq y+\Delta | T > y)}{\Delta}$$

$$P(A|B) = \frac{P(A \cap B)}{P(B)}$$

That's why this is ^{like a} conditional density. But

it's not a real conditional density since it does not integrate to 1.

If $T \sim \text{Exp}(\lambda)$

$$h(y) = \frac{\lambda e^{-\lambda y}}{\underbrace{e^{-\lambda y}}_{P(T > y)}} = \lambda \quad (\text{constant})$$

Thus, attach iid edge weights to the lattice $\text{Exp}(\lambda)$ and at the time at which one of the exponential clocks on the edges fires, grow the cluster to that edge.

$$X = \min\{T_1, \dots, T_u\} \quad \text{Suppose } T_k \neq \arg\min\{T_1, \dots, T_u\} \\ =: i_0$$

$$P(T_k > X+u \mid k \neq i_0) = P(T_k > X+u \mid T_k > X)$$

$$= \int_0^{\infty} P(T_k > t+u \mid T_k > t) \lambda e^{-\lambda t} dt$$

$$= \int_0^{\infty} \frac{e^{-\lambda(t+u)}}{e^{-\lambda t}} \lambda e^{-\lambda t} dt = e^{-\lambda u}$$

So by the memoryless property, we can just replace all the T_k everytime a vertex is turned red by new independent \tilde{T}_k . This is called 'restarting'

the exponential clocks.

So here's the punchline: The Eden model is just FPP with exponential edge wts, but we IGNORE the times at which vertices are added to the growing cluster.

Richardson proved the shape theorem (in mean)

$$P\left((1-\epsilon)B \subset \mathcal{O}_n \subset (1+\epsilon)B_n\right) \xrightarrow{n \rightarrow \infty} 1$$

Eden's conjecture is kind of boring, but Richardson proved it.


SKIP
(Eden's conjecture says that the fraction of configurations with n cells C_n likely to be generated tends to 0.)

$C_n = \#$ of configs with n cells containing the origin.

$\geq 2^n$ (random walk that goes up or right)

The config with n cells that is actually generated will be

contained inside two balls of radius $(k-\epsilon)\sqrt{n}$ and

$(k+\epsilon)\sqrt{n}$.  # of cells = $2\pi k\sqrt{n}\epsilon$
Total # of configs $\leq 2^{2\pi k\sqrt{n}\epsilon}$

Thus
$$\frac{2\pi k\sqrt{n}\epsilon}{2^n} \rightarrow 0$$

On page 13, he states:

" As p increases, the roughness of the boundary seems to increase. As the roughness of the boundary increases, the natural norm imposed by the lattice breaks down, and the Euclidean norm asserts itself. It would be interesting to have an explanation for this phenomenon.

He only goes to $t=50$ (He does geometric FPP)

Richardson's conjecture is still open (Is Exponential GDP a ball?)

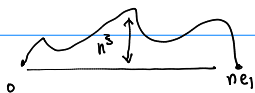
In the meantime, in the physics world

Kardar - Parisi - Zhang introduce KPZ scaling etc (1986)
Phy Rev. Lett.

Zabolitsky - Stauffer simulate the Eden model (1986)
Phys. Rev. A.

The general conjecture was that in 2D surface growth

$$\sqrt{\text{Var}(T(0, ne_1))} \approx n^\chi \leftarrow \begin{array}{l} \text{fluctuation} \\ \text{exponent} \end{array}$$



$$\chi = \frac{1}{3} \text{ and } \bar{\xi} = \frac{2}{3} \text{ ind}=2$$

$$\underbrace{2\bar{\xi} - 1 = \chi}_{\text{KPZ relationship.}} \text{ in general } d$$

The 90s saw a lot of mathematical activity (independent of physics) but no simulation studies had been performed.

This gap was filled finally in 2013 (Ahn & Deijfen)

But even now their simulations seem small compared to what one could do on a modern ^{set of} GPU.

↳ Could ask Sneepathi Pai at OR if this is a possibility.

Their main question is the following:

Given cdfs F and \tilde{F} , and corresponding FPP models.

When is $M_{\tilde{F}} \leq M_F$?

We know vBK criterion. If $F(0) < p_c$ and

$$F <_{\text{var}} \tilde{F} \quad \left(\int \phi d\tilde{F} \leq \int \phi dF \quad \forall \text{ concave increasing } \phi \right)$$

then $M_F < M_{\tilde{F}}$ (if $F \neq \tilde{F}$)

In particular if

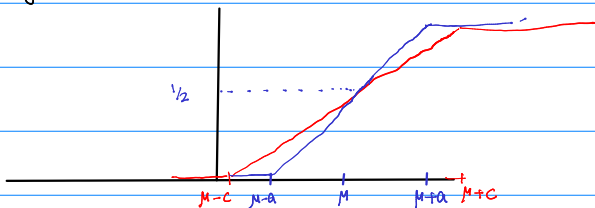
$$F = \text{Unif} [M-a, M+a]$$

$$\tilde{F} = \text{Unif} [M-c, M+d]$$

such that

$$\left[\begin{array}{cc} \mu & \\ \mu-c & \mu-a \end{array} \right] \left[\begin{array}{cc} \mu & \\ \mu+a & \mu+c \end{array} \right]$$

Then by the cut criterion, \tilde{F} is more variable than F



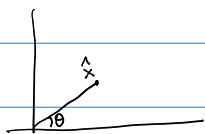
In this case, as variance increases, then so does the time constant. Is this true in general?

Alm and Deijfen propose a new quantity:

$$E[\min\{t_1, t_2, t_3, t_4\}] =: E_4[t]$$

that appears to "determine" M_F .

This is how they setup their simulations:



\hat{x} unit vector with Euclidean norm in direction θ .

They choose $\theta = 0, 14, 27, 37, 45^\circ$ and find M by growing

the cluster until it hits a line.

They pick several different distributions:

1) $Unif(0, 1)$ 3) $\Gamma(2, 2)$ Gamma distribution

2) $Exp(1)$ 4) $\Gamma(3, 3)$, $\Gamma(4, 4)$

5) $U(0.1, 0.9)$ 5) $\hat{F}_i(k)$ $k = 1, 2, 3, 4$

Weird outlier

$$Tail(x) = \frac{1}{(1+x)^k}$$

(Previously Eden, Richardson had focussed only on the Exponential case.)

I'll return to this paper many times. For example, it has 1) Non random fluctuations (Fig 2)] Skip

2) A linear graph relating $M(e_i) = 1.5 E_u[Z]$

and $M(\hat{X}_{450}) = 1.56 E_u[Z]$ (Fig 4 and 5 on page 13)

3) 2 distributions F with the same mean

and $\text{Var}(F) = \text{Var}(\hat{F})$ but $M_{\hat{F}} < M_F$

"Proving" that variance alone cannot successfully describe the time constant (Fig 3)

4) An observation that $M(\hat{X}_\theta)$ is increasing

on the interval $\theta \in [0, \frac{\pi}{4}]$ (It's enough to describe

M on an octant and use symmetry)

Research Problem: Prove this. I think this might be

quite hard to do, BUT I think a more sensible

norm here is the $|\cdot|_1$ norm. I wonder if the monotonicity

is still observed in this norm.

5) They also observe that the "most heavy

tailed" distributions are the most circle

like (Fischer distributions on page 16,

Fig 9)

6) In Fig 10 they simulate a bunch of $\text{Exp}(1) + c$ distributions and show that the shape gets more and more like a diamond.

Ex: Prove this by considering $\frac{\text{Exp}(1)+c}{c}$ and applying Cox - Kerstan.

7) Then they show that $\chi = 0.322 \dots$ $\xi = 0.667$ in their experiments (Fig 16)

8) Alm and Deijfen compute $T(0, L_n)$ instead of $T(0, n_1)$

(this works for time constant computation due to a theorem of Wiernman and Ren). The conjecture that

$$\text{Var}(T(0, L_n)) \leq \text{Var}(T(0, n_1)).$$

Can one prove this?

8) Finally, let $E_4(z^2) := E \left[\min_{i=1, \dots, 4} z_i^2 \right]$

In $\hat{A}(0.3)$ with tail $\frac{1}{(1+x)^2}$ we have $E_4[z] < \infty$
but $E_4(z^2) = +\infty$.

In this case, Richardson's theorem is not true.

$$\mathbb{P} \left((1-\epsilon) \mathbb{B} \subset \frac{\mathbb{B}_n}{n} \subset (1+\epsilon) \mathbb{B} \right) \rightarrow 1$$

This is seen in Fig 15. The boundary of the infected set is shown in red. The red regions in the interior enclose regions that have not yet been

infected. This is the content of the next lecture.

The Cox - Durrett limit shape theorem.

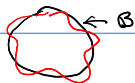
Recall p_c , the threshold for Bernoulli bond percolation.

Suppose $\left[\begin{array}{l} \mathbb{E} \left[\min_{i=1, \dots, d} \{t_i\} \right] < \infty \quad \left(\begin{array}{l} \text{a bit more than} \\ \text{time constant} \\ \text{condition} \end{array} \right) \text{---} \textcircled{\#1} \\ F(0) < p_c \quad \left(\begin{array}{l} \text{no percolation of} \\ 0s \Rightarrow H(\epsilon) \neq 0 \end{array} \right) \text{---} \textcircled{\#2} \end{array} \right.$

GOOD measures

Theorem : (Cox and Durrett)

Let $B = \{x \mid \mu(x) \leq 1\}$



Let $B(n) = \{x \in \mathbb{R}^d \mid T(0, x) \leq n\}$

Then for any $\epsilon > 0$

$$P\left((1-\epsilon)B \subset \frac{B(n)}{n} \subset (1+\epsilon)B \text{ for all large } n \right) = 1$$

Alm and Deift also have many intriguing questions : Is $\mu \geq E_4[Z]$ in general?

This would be LOVELY to prove.